# Simple linear regression (SLR) 

Anna Helga Jonsdottir<br>Gunnar Stefansson<br>Sigrun Helga Lund

University of Iceland

## Introduction

The following gives a review of:

- Scatter plots
- Correlation
- Simple linear regression - SLR
- Inference in SLR


## Scatter plot

## Scatter plot

Scatter plots are used to investigate the relationship between two numerical variables.

The value of one variable is on the $y$-axis (vertical) and the other on the $x$-axis (horizontal).

When one of the variable is an explanatory variable and the other one is a response variable, the response variable is always on the $y$-axis and the explanatory variable on the $x$-axis.


## The equation of a straight line

The equation of a straight line
The equation of a straight line describes a linear relationship between two variables, $x$ and $y$. The equation is written

$$
y=\beta_{0}+\beta_{1} x
$$

where $\beta_{0}$ is the intercept of the line on the $y$-axis and $\beta_{1}$ is the slope of the line.

## The straight line



## Linear relationship

## Linear relationship

We say that the relationship between two variables is linear if the equation of a straight line can be used to predict which value the response variable will take based on the value of the explanatory variable.

There can be all sorts of relationship between two variables. For example, the relationship can be described with a parabola, an exponential function and so on. Those relationship are referred to as nonlinear.

## Linear/nonlinear relationship



Nonlinear relationship


Linear relationship


Nonlinear relationship


## Sample correlation coefficient

Sample coefficient of correlation
Assume that we have $n$ measurements on two variables $x$ and $y$. Denote the mean and the standard deviation of the variable $x$ with $\bar{x}$ and $s_{x}$ and the mean and the standard deviation of the $y$ variable with $\bar{y}$ and $s_{y}$. The sample coefficient of correlation is

$$
r=\frac{1}{n-1} \sum_{i=1}^{n}\left(\frac{x_{i}-\bar{x}}{s_{x}}\right)\left(\frac{y_{i}-\bar{y}}{s_{y}}\right)
$$

Warning: The correlation only estimates the strength of a linear relationship!

Perfect positive relationship


No relationship




Perfect
negative relationship


No relationship


Weaker positive relationship



## Correlation and causation

- Causation is when changes in one variable cause changes in the other variable.
- There is often strong correlation between two variables although there is no causal relationship.
- In many cases, the variables are both influenced by the third variable which is then a lurking variable.
- Therefore, high correlation on its own is never enough to claim that there is a causal relationship between two variables!


## Informal regression

Input: Have data as $(x, y)$-pairs
Suppose a scatterplot indicates a linear relationship
Loosely: Want to "fit a line" through the data
Next: Evaluate the fit

## Formal regression

Consider fixed numbers, $x_{i}$
Random variables: $Y_{i} \sim N\left(\beta_{0}+\beta_{1} x_{i}, \sigma^{2}\right)$
or: $Y_{i}=\beta_{0}+\beta_{1} x_{i}+\epsilon_{i}$
where $\epsilon_{i} \sim N\left(0, \sigma^{2}\right)$ is a random error term, independent and identically distributed (i.i.d.)

We collect some data on $\left(y_{i}, x_{i}\right)$ and use the data to estimate $\beta_{0}$ and $\beta_{1}$. We can then predict $Y$ using

$$
\hat{y}_{i}=b_{0}+b_{1} x_{i}
$$

Then

$$
e_{i}=y_{i}-\hat{y}_{i}
$$

is the $i$ th residual - the difference between the $i$ th observed response and the prediction.


## The linear regression model

The linear regression model
The simple linear regression model is written

$$
Y=\beta_{0}+\beta_{1} x+\varepsilon
$$

where $\beta_{0}$ and $\beta_{1}$ are unknown parameters and $\varepsilon$ is a normally distributed random variable with mean 0 and variance $\sigma^{2}$.

The aim of the simple linear regression is first and foremost to estimate the parameters $\beta_{0}$ and $\beta_{1}$ with the measurements on the two variables, $x$ and $Y$.

The most common estimation method is through least squares.

## Which line?



## The least squares method



## The least squares method

Least squares estimation technique minimizes:

$$
S=\sum\left(y_{i}-\left(b_{0}+b_{1} x_{i}\right)\right)^{2}
$$

Maximum likelihood estimation assumes a probability distribution for the data and maximizes the corresponding likelihood function. In the case of normal distributions the two methods results in the same estimates - we will use least squares.

## The least squares regression line

Let the mean and standard deviation of the $x$ variable with $\bar{x}$ and $s_{x}$ and the $y$ variable with $\bar{y}$ and $s_{y}$ and their correlation coefficient with $r$.
Let $b_{0}$ denote the estimate of $\beta_{0}$ and $b_{1}$ denote the estimate of $\beta_{1}$. Then $b_{0}$ and $b_{1}$ are given with the equation

$$
b_{1}=\frac{\sum(x-\bar{x})(y-\bar{y})}{\sum(x-\bar{x})^{2}}=r \frac{s_{y}}{s_{x}}
$$

and

$$
b_{0}=\bar{y}-b_{1} \bar{x} .
$$

These are the least squares estimates of the coefficient of a regression line through the data points $(x, y)$.

## Prediction

We often want to use our regression model to predict the outcome of our response variable for some value(s) of the explanatory variable.

Prediction
We can predict the value of $Y$ for some value of $x$ using

$$
\hat{y}_{h}=b_{0}+b_{1} \cdot x_{h}
$$

## Interpolation

## Interpolation

If the regression model is used to predict a value of $Y$ for some value of $x$ which is similar to the $x$-values that were used to estimate the model is referred to as interpolating.

## Extrapolation

## Extrapolation

Extrapolating is using the regression model to predict a value of $Y$ for some value of $x$ which is far from the $x$-values that were used to estimate the model.

It can be very questionable to extrapolate!

## Estimating dispersion

A point estimate of $\sigma^{2}$, the variance of the $y$-measurements, is obtained with

$$
s^{2}=\frac{\sum_{i}\left(y_{i}-\left(b_{0}+b_{1} x_{i}\right)\right)^{2}}{n-2}
$$

The predicted value of $y$ at a given $x$ is often denoted by $\hat{y}=b_{0}+b_{1} x$ and therefore

$$
s^{2}=\frac{\sum_{i}\left(y_{i}-\hat{y}_{i}\right)^{2}}{n-2}
$$

Commonly $\hat{\sigma}^{2}$ is used in place of $s^{2}$.

## Estimating dispersion

We will commonly use the notation

$$
S S_{E}=\sum_{i}\left(y_{i}-\left(b_{0}+b_{1} x_{i}\right)\right)^{2}
$$

and

$$
M S E=S S_{E} /(n-2)
$$

so $s^{2}=M S E$.

## Correlation and explained variation

Recall the the correlation coeffficient $r$ is always between -1 and 1 .
Write $S S_{E}=\sum(y-\hat{y})^{2}$ (sum of squared errors, i.e. error after regression), and $S S_{T O T}=\sum(y-\bar{y})^{2}$ (total sum of squares, i.e. before regression)

The explained variation
The explained variation, often called the coefficient of determination, is calculated with

$$
R^{2}=1-\frac{S S_{E}}{S S_{T O T}}
$$

Note:

$$
R^{2}=1-\frac{\sum(y-\hat{y})^{2}}{\sum(y-\bar{y})^{2}}=\ldots=r^{2}
$$

## SLR in $R$

It is easy to perform linear regression in R using the $\operatorname{lm}()$ function. For simple linear regression the syntax is
fit <- lm(x ~ y, data=nameofdataset)

The results can then be looked at using the summary () function summary (fit)

## Inference in the linear regression model

Recall that if we have $n$ paired measurements $\left(x_{1}, y_{1}\right), \ldots,\left(x_{n}, y_{n}\right)$, the regression model can be written as

$$
Y_{i}=\beta_{0}+\beta_{1} x_{i}+\varepsilon_{i} .
$$

- $\beta_{0}$ is the true intercept (population intercept) that we do not know the value of.
- $\beta_{1}$ is the true slope (population slope)
- $\epsilon_{i}$ are the errors where $\epsilon \sim N\left(0, \sigma^{2}\right)$ and

$$
\hat{\sigma}^{2}=\frac{\sum(y-\hat{y})^{2}}{n-2}
$$

$\beta_{0}$ and $\beta_{1}$ are therefore parameters, that we both want to estimate and make inference on.

## Estimating slope and intercept accuracy

The standard error of the slope is:

$$
\hat{\sigma}_{b_{1}}^{2}=\frac{\hat{\sigma}^{2}}{\sum(x-\bar{x})^{2}}
$$

and the standard error of the intercept is:

$$
\hat{\sigma}_{b_{0}}^{2}=\left(\frac{1}{n}+\frac{\bar{x}^{2}}{\sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}}\right) \hat{\sigma}^{2}
$$

where

$$
\hat{\sigma}^{2}=\frac{\sum(y-\hat{y})^{2}}{n-2}
$$

## Elements of inference in simple linear regression

- Basic inference: Test hypotheses and generate confidence intervals for slope and intercept.
- Results on the estimators can be used to make inference on the true slope and intercept.
- The first question raised is whether there is any relationship between the $x$ and $y$ measurements, i.e. whether the slope is zero. This can be phrased as a general hypothesis test for the slope.
- Although hypothesis tests are important, they give no information if the hypothesis can not be rejected and hence confidence intervals tend to be more informative in general.
- Both hypothesis tests and confidence intervals can be derived for the intercept as well as the slope, although inference for the intercept tends not to be as commonly used.


## Testing hypotheses concerning the slope

Hypothesis test for $\beta_{1}$
The null hypothesis is:

$$
H_{0}: \beta_{1}=\beta_{1_{0}}
$$

The test statistic is:

$$
t=\frac{b_{1}-\beta_{1_{0}}}{\hat{\sigma}_{\hat{\beta}_{1}}}
$$

If the null hypothesis is true the test statistic follows the t distribution with $\mathrm{n}-2$ degrees of freedom or $t \sim t(n-2)$.

Alternative hypothesis Reject $H_{0}$ if:

| $\mathrm{H}_{1}: \beta_{1}<\beta_{1_{0}}$ | $t<-t_{1-\alpha}$ |
| :---: | :---: |
| $\mathrm{H}_{1}: \beta_{1}>\beta_{1_{0}}$ | $t>t_{1-\alpha}$ |
| $\mathrm{H}_{1}: \beta_{1} \neq \beta_{1_{0}}$ | $t<-t_{1-\alpha / 2}$ or $t>t_{\alpha / 2}$ |

## Confidence interval for $\beta_{1}$

Confidence interval for $\beta_{1}$
The lower bound of $1-\alpha$ confidence interval for $\beta_{1}$ is:

$$
b_{1}-t_{1-\alpha / 2,(n-2)} \cdot \hat{\sigma}_{\hat{\beta}_{1}}
$$

The upper bound of $1-\alpha$ confidence interval is:

$$
b_{1}+t_{1-\alpha / 2,(n-2)} \cdot \hat{\sigma}_{\hat{\beta}_{1}}
$$

where $b_{1}$ is calculated the same way as usual, $n$ is the number of paired measurements and $t_{1-\alpha / 2,(n-2)}$ is found in the t -distribution table.

## Confidence interval for $\beta_{0}$

Confidence interval for $\beta_{0}$
The lower bound of a $1-\alpha$ confidence interval for $\beta_{0}$ is:

$$
b_{0}-t_{1-\alpha / 2,(n-2)} \cdot \hat{\sigma}_{\hat{\beta}_{0}}
$$

The upper bound of $1-\alpha$ confidence interval is:

$$
b_{0}+t_{1-\alpha / 2,(n-2)} \cdot \hat{\sigma}_{\hat{\beta}_{0}}
$$

where $b_{0}$ is calculated the same way as usual, $n$ is the number of paired measurements and $t_{1-\alpha / 2,(n-2)}$ is in the table for the t -distribution.

## Confidence interval for a point on the regression line

Confidence interval for a point on the regression line
The lower bound of $1-\alpha$ confidence interval for $\hat{Y}_{h}$ is:

$$
\left(b_{0}+b_{1} x_{h}\right)-t_{1-\alpha / 2,(n-2)} \cdot s_{y_{h}}
$$

The upper bound of $1-\alpha$ confidence interval is:

$$
\left(b_{0}+b_{1} x_{h}\right)+t_{1-\alpha / 2,(n-2)} \cdot s_{y_{h}}
$$

where $b_{0}$ and $b_{1}$ are calculated the same way as usual, $n$ is the number of paired measurements, $t_{1-\alpha / 2,(n-2)}$ is found in the t-distribution table and

$$
s_{y_{h}}=\sqrt{\hat{\sigma}^{2}\left(\frac{1}{n}+\frac{\left(x_{h}-\bar{x}\right)^{2}}{\sum_{j}\left(x_{j}-\bar{x}\right)^{2}}\right)}
$$

## Predicting a new observation

Notice, that as a prediction for a future point, includes two sources of variation or error, first due to the measurement errors in the original data through variation in the parameter estimates and secondly through the future measurement errors at this point.

## Prediction interval for a new observation

Prediction interval for a new observation
The lower bound of $1-\alpha$ prediction interval for $\hat{Y}_{h}$ is:

$$
\left(b_{0}+b_{1} x_{h}\right)-t_{1-\alpha / 2,(n-2)} \cdot s_{\text {pred }}
$$

The upper bound of $1-\alpha$ prediction interval is:

$$
\left(b_{0}+b_{1} x_{h}\right)+t_{1-\alpha / 2,(n-2)} \cdot s_{\text {pred }}
$$

where $b_{0}$ and $b_{1}$ are calculated the same way as usual, $n$ is the number of paired measurements, $t_{1-\alpha / 2,(n-2)}$ is found in the t-distribution table and

$$
s_{\text {pred }}=\sqrt{\hat{\sigma}^{2}\left(1+\frac{1}{n}+\frac{\left(x_{h}-\bar{x}\right)^{2}}{\sum_{j}\left(x_{j}-\bar{x}\right)^{2}}\right)} .
$$

